Trends in the Timing of Pre-emptive Kidney Transplantation

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ABSTRACT

Pre-emptive kidney transplantation is considered the best available renal replacement therapy, but no guidelines exist to direct its timing during CKD progression. We used a national cohort of 19,471 first-time pre-emptive kidney transplant recipients between 1995–2009 to evaluate patterns and implications of transplant timing. Mean estimated GFR (eGFR) at the time of pre-emptive transplant increased significantly over time, from 9.2 ml/min/1.73m² in 1995 to 13.8 ml/min/1.73m² in 2009 (P<0.001). Patients with eGFR≥15 ml/min/1.73m² represented an increasing proportion of pre-emptive transplant recipients, from 9% in 1995 to 35% in 2009; the trend for patients with eGFR≥10 was similar (30% to 72%). We did not detect statistically significant differences in patient survival or death-censored graft survival between strata of eGFR at the time of transplant, either in the full cohort or in subgroup analyses of patients who might theoretically benefit from earlier pre-emptive transplantation. In summary, pre-emptive kidney transplantation is occurring at increasing levels of native kidney function. Earlier transplantation does not appear to associate with patient or graft survival, suggesting that earlier pre-emptive transplantation may subject donors and recipients to premature operative risk and waste the native kidney function of recipients.

“early PKT”) increased from 9.2% in 1995 to 34.7% in 2009; the proportion with eGFR/H11350 increased from 30.0% to 72.4%. Trends were comparable within deceased donor and living donor PKT recipients. Among patients undergoing PKT, those transplanted more recently, those with diabetes or tubulointerstitial diseases, and those with poor functional status were more likely to do so at eGFR/H11350 (adjusted relative rate [aRR] 1.12, 1.52, 1.30, and 1.14, respectively; Table 1). Patients with PKD, women, African Americans, and other nonwhites were less likely to undergo early PKT (aRR 0.77, 0.87, 0.90, and 0.78, respectively). Age was negatively associated with early PKT (9% less likely to undergo early PKT with every increasing decade) until the age of 50, after which there was no significant association. The association between some of these factors and early PKT changed significantly over time, as measured by interaction term analysis. Over time, diabetics became less likely to receive early PKT (6% smaller yearly increase than nondiabetics, \( P = 0.001 \)), whereas patients with PKD became more likely (3% greater yearly increase than non-PKD patients, \( P = 0.023 \)). Similarly, a greater proportion of women underwent early PKT in recent years (3% greater yearly increase than men, \( P = 0.003 \)).

There were no differences in unadjusted patient or graft survival by pre-PKT eGFR (Figure 2), and a possible trend toward worse propensity score-adjusted patient survival associated with higher pretransplant eGFR (Table 2). Results comparing early (eGFR≥15) with later PKT were similar by various methods of accounting for confounding and treatment selection bias: propensity score-adjusted analyses (adjusted hazard ratio [aHR] for death, 1.10, 95% CI 0.98 to 1.24, \( P = 0.1 \); aHR for death-censored graft loss, 0.96, 95% CI 0.84 to 1.08, \( P = 0.5 \)); multivariate Cox proportional hazards analyses (live donors: aHR for death, 0.89, 95% CI 0.74 to 1.08, \( P = 0.2 \); aHR for graft loss, 1.01, 95% CI 0.84 to 1.21, \( P = 0.9 \)); deceased donors: aHR for death, 1.16, 95% CI 1.00 to 1.35, \( P = 0.1 \); aHR for graft loss, 0.99, 95% CI 0.83 to 1.19, \( P = 0.9 \)), and instrumental variable analyses (adjusted odds ratio [aOR] for death, 1.01, 95% CI 0.91 to 1.13, \( P = 0.8 \); aOR for graft loss 0.98, 95% CI 0.87 to 1.11, \( P = 0.8 \)). Similarly, no statistically significant differences were seen when limiting the analyses to subgroups of recipients that might theoretically benefit from early PKT (diabetics, recipients over 65, those with poor functional status, and men; Supplemental Figure 1).

This large national observational study identifies a dramatic trend toward PKT earlier in the course of CKD. Among PKT recipients, patients receiving PKT at eGFR≥15 were more often Caucasian, male, diabetic, with worse functional status. However, even after adjustment for these differences and potential treatment selection bias, early PKT was not correlated with improved patient or graft survival. Early PKT may introduce risks of transplantation and “start the clock” on allograft survival unnecessarily early.

The trend toward early PKT contrasts with published suggestions that late PKT may be preferable. An older study of 4046 PKT recipients found no associa-

Figure 1. Recipient native eGFR at PKT. (A) Distributions of eGFR in 1995, 2002, and 2009 are illustrated by kernel density plots. (B) Distribution of eGFR categories among PKT recipients, 1995–2009.
tion between pretransplant eGFR ≥15 and PKT outcomes, including graft survival, patient survival, and 6-mo posttransplant eGFR. A two-center study of 671 PKT recipients looked at additional categories of pretransplant eGFR and found no benefit with earlier PKT. In our study, using the largest cohort of PKT recipients to date, spanning a 15-yr time period, and using multiple analytic methods to minimize bias, we were also unable to demonstrate a benefit of early PKT, even in populations that intuitively might benefit from early restoration of renal function (e.g., those with a greater burden of comorbid disease or a higher risk of cardiovascular events).

Given the consistent lack of benefit, the rationale driving earlier PKT is unclear. It may be an unintended consequence of the push to "transplant first." Estimating rates of kidney function decline is difficult, and current estimates may stem from historical rates that are now inaccurate given improvements in CKD care. Earlier PKT may be pursued to avoid unanticipated acceleration of CKD progression, dialysis requirement, and the morbidity associated with vascular access creation. Nationally, dialysis initiation is occurring earlier, which, in turn, forces earlier PKT. Alternatively, transplant timing may be unduly driven by donor factors. Approving a suitable living donor can be a lengthy process, and programs may be reluctant to delay transplantation once a donor is cleared. Early PKT may be pursued out of a concern that a donor may become unable to donate, or to accommodate a donor’s schedule.

The implications of early PKT may be worrisome. The parallel trend in early-start dialysis has been associated with worse or no different outcomes when compared with later dialysis initiation. Early intervention, whether dialysis or transplantation, increases costs and significantly changes a patient’s life, with required adherence to new medication or treatments. In living donor PKT, both recipient and donor may undergo potentially premature operative risk. Given the enormity of the organ shortage and the independence of pretransplant eGFR and graft survival, premature transplantation may waste allograft life of a precious resource.

There are several important limitations in this study. First and foremost, the data are observational. There may be treatment selection bias, where recipients undergo early PKT for reasons correlated with survival. For example, sicker patients may be referred more often for early PKT. However, we used several analytic methods designed to minimize this bias, including propensity score adjustment and instrumental variable analysis, an approach that may more closely approximate results of randomized controlled trials. Furthermore, results are similar within deceased donor PKT recipients, a group in which treatment selection bias should be reduced because of unpredictable organ availability.

Second, starting the “survival clock” at PKT requires that a patient survive and remain off dialysis until transplantation. As a result, the late PKT group may be healthier; They were able to survive longer CKD progression without development of uremia. Additionally, measuring survival from the time of PKT does not credit late PKT recipients with this period of CKD progression, nor does it account for those who would-be late PKT recipients who died before transplant. A better method to compare early versus late PKT timing might be to measure survival from the time potential PKT recipients reach a given eGFR, although this would not be possible using transplant registry data.

Finally, kidney function at PKT was determined using only one value of non-standardized serum creatinine, and we assumed the patient was in steady-state. More accurate estimations of native kid-
ney function, such as measured GFR, may improve estimations of relative risk.

In summary, we report that PKT is being performed earlier, with more patients transplanted at higher levels of native kidney function. There was no demonstrated survival advantage to PKT at higher levels of eGFR, with results robust to three methods of accounting for potential bias in observational studies. Until better data are available, providers and transplant programs should strive to individualize transplant timing decisions to avoid PKT too early in the course of CKD.

**CONCISE METHODS**

**Study Population**
A national cohort of 152,731 first-time, kidney-only adult transplant recipients between January 1, 1995, and December 31, 2009, was identified through the UNOS registry. PKT recipients (n = 19,471) were defined by the absence of exposure to dialysis before kidney transplantation. Ascertainment of death was supplemented by linkage to the Social Security Death Master File.

**Definitions**
Estimated GFR (eGFR) was calculated using the four-variable Modification of Diet in Renal Disease (MDRD) equation, using serum creatinine at the time of transplantation as supplied to UNOS. Pretransplant serum creatinine was available for 95.9% of living donor PKT recipients and 92.4% of deceased donor PKT recipients over the study period. Early PKT was defined as PKT performed when native eGFR was ≥15 ml/min/1.73 m².

**Statistical Analysis**
Associations with early PKT were estimated using a generalized linear model, as described previously. To determine if the early PKT trends were different among different patient phenotypes, interactions with year of transplant were tested with five factors specified a priori: age, gender, cause of ESRD, donor type, and functional status.

**Accounting for Potential Bias**
Outcomes associated with early PKT were estimated in three ways: propensity score-adjusted Cox proportional hazards regression, multivariate Cox proportional hazards regression, and instrumental variable analysis. Propensity scores were derived using random forest regression, a machine learning regression tree technique designed to account for all possible levels of multi-interaction; individual propensity scores for early PKT were categorized into quartiles and included in a survival model of all PKT recipients. To test the sensitivity of our inferences to the method of propensity score derivation, the analysis was repeated using propensity scores derived from simple multivariate logistic regression, and no inferential difference was noted. Covariates in the multivariate survival analyses included transplant year as well as all vari-

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**Table 2. PKT recipient and graft survival associated with pretransplant eGFR**

<table>
<thead>
<tr>
<th>eGFR at PKT</th>
<th>Adjusted HR of Death</th>
<th>Adjusted HR of Death-Censored Graft Loss</th>
</tr>
</thead>
<tbody>
<tr>
<td>eGFR &lt;10</td>
<td>reference</td>
<td>reference</td>
</tr>
<tr>
<td>eGFR 10–15</td>
<td>1.10 (95% CI 0.99–1.21, P = 0.07)</td>
<td>0.97 (95% CI 0.88–1.08, P = 0.0)</td>
</tr>
<tr>
<td>eGFR 15–20</td>
<td>1.16 (95% CI 1.00–1.34, P = 0.05)</td>
<td>0.95 (95% CI 0.81–1.11, P = 0.5)</td>
</tr>
<tr>
<td>eGFR ≥20</td>
<td>1.12 (95% CI 0.93–1.34, P = 0.2)</td>
<td>0.94 (95% CI 0.77–1.15, P = 0.5)</td>
</tr>
</tbody>
</table>

*Propensity score-adjusted.

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**Figure 2.** Kaplan-Meier estimates of (A) patient survival, and (B) death-censored graft survival in PKT recipients, by pretransplant eGFR.
ables specified in standard Scientific Registry for Transplant Recipients (SRTR) outcome models for 3-yr postkidney transplant patient survival; the model was specified separately for deceased donor and living donor recipients to remain consistent with the SRTR models. Finally, instrumental variable analysis was performed for the binary outcomes of 3-yr post-PKT mortality and death-certified graft loss, using regional rate of early PKT as the instrumental variable and adjusting for covariates listed in Table 1. All analyses were performed using STATA 11.0/MP for Linux (College Station, TX) except for the random forest regression, which was implemented using R 2.7.2 (The R Foundation for Statistical Computing, Vienna, Austria).

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DISCLOSURES

None.

REFERENCES


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